

Challenges in Fiducial Inference

Parts of this talk are joint work with

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^aNSF support acknowledged

Outline

- Introduction
- Definition
- Sparsity
- Regularization
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Fiducial?

- ▶ **Oxford English Dictionary**

- ▶ adjective technical (of a point or line) used as a fixed basis of comparison.
- ▶ Origin from Latin fiducia 'trust, confidence'

- ▶ **Merriam-Webster dictionary**

1. taken as standard of reference *a fiducial mark*
2. founded on faith or trust
3. having the nature of a trust : fiduciary

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- ▶ Challenge of extra information:
 - ▶ Sparsity
 - ▶ Regularization
- ▶ My point of view: frequentist
 - ▶ Justified using asymptotic theorems and simulations.
 - ▶ GFI tends to work well

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- ▶ **Likelihood** is the function $f(\mathbf{x}, \xi)$, where ξ is variable and \mathbf{x} is fixed.
 - ▶ Likelihood as a distribution?

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$$\arg \min_{\xi} \|\mathbf{x} - \mathbf{G}(\mathbf{U}^*, \xi)\| \mid \left\{ \min_{\xi} \|\mathbf{x} - \mathbf{G}(\mathbf{U}^*, \xi)\| \leq \varepsilon \right\} \quad (1)$$

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- ▶ Similar to ABC; generating from prior replaced by **min**.
- ▶ Is this practicle? Can we compute?

Explicit limit (1)

- ▶ Assume $\mathbf{X} \in \mathbb{R}^n$ is continuous; parameter $\xi \in \mathbb{R}^p$
- ▶ The limit in (1) has density (H, Iyer, Lai & Lee, 2016)

$$r(\xi|\mathbf{x}) = \frac{f_{\mathbf{X}}(\mathbf{x}|\xi)J(\mathbf{x}, \xi)}{\int_{\Xi} f_{\mathbf{X}}(\mathbf{x}|\xi')J(\mathbf{x}, \xi') d\xi'},$$

where $J(\mathbf{x}, \xi) = D \left(\left. \frac{d}{d\xi} \mathbf{G}(\mathbf{u}, \xi) \right|_{\mathbf{u}=\mathbf{G}^{-1}(\mathbf{x}, \xi)} \right)$

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- ▶ $n = p$ gives $D(A) = |\det A|$
- ▶ $\|\cdot\|_2$ gives $D(A) = (\det A^T A)^{1/2}$
Compare to Fraser, Reid, Marras & Yi (2010)
- ▶ $\|\cdot\|_\infty$ gives $D(A) = \sum_{\mathbf{i}=(i_1, \dots, i_p)} |\det(A)_{\mathbf{i}}|$

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$$\begin{aligned} J(\mathbf{y}, \beta, \sigma) &= \left(\det \left(\left(\mathbf{X}, \frac{\mathbf{y} - \mathbf{X}\beta}{\sigma} \right)^\top \left(\mathbf{X}, \frac{\mathbf{y} - \mathbf{X}\beta}{\sigma} \right) \right) \right)^{1/2} \\ &= \sigma^{-1} |\det(\mathbf{X}^T \mathbf{X})|^{1/2} (RSS)^{1/2} \end{aligned}$$

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- ▶ Fiducial *happens to be* same as independence Jeffreys, *explicit* normalizing constant

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 - ▶ In simulations fiducial was marginally better than reference prior which was much better than flat prior.

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- ▶ GFD is *not* invariant to smooth transformation of the data if $n > p$
- ▶ Does not satisfy likelihood principle.

Various Asymptotic Results

$$r(\xi|\mathbf{x}) \propto f_{\mathbf{X}}(\mathbf{x}|\xi)J(\mathbf{x}, \xi) \text{ where } J(\mathbf{x}, \xi) = D \left(\left. \frac{d}{d\xi} \mathbf{G}(\mathbf{u}, \xi) \right|_{\mathbf{u}=\mathbf{G}^{-1}(\mathbf{x}, \xi)} \right)$$

- ▶ Most start with $C_n^{-1}J(\mathbf{x}, \xi) \rightarrow J(\xi_0, \xi)$
- ▶ Bernstein-von Mises theorem for fiducial distributions provides asymptotic correctness of fiducial CIs H (2009, 2013), Sonderegger & H (2013) .
- ▶ Consistency of model selection H & Lee (2009), Lai, H & Lee (2015), H, Iyer, Lai & Lee (2016).
- ▶ Regular higher order asymptotics in Pal Majumdar & H (2016+).

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Model Selection

$$\blacktriangleright \mathbf{X} = \mathbf{G}(M, \boldsymbol{\xi}_M, \mathbf{U}), \quad M \in \mathcal{M}, \boldsymbol{\xi}_M \in \boldsymbol{\xi}_M$$

Theorem: (H, Iyer, Lai, Lee 2016) Under assumptions

$$r(M|\mathbf{y}) \propto q^{|\mathcal{M}|} \int_{\boldsymbol{\xi}_M} f_M(\mathbf{y}, \boldsymbol{\xi}_M) J_M(\mathbf{y}, \boldsymbol{\xi}_M) d\boldsymbol{\xi}_M$$

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- ▶ Need for penalty – in fiducial framework additional equations

$$0 = P_k, \quad k = 1, \dots, \min(|M|, n)$$
 - ▶ Default value $q = n^{-1/2}$ (motivated by MDL)

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$$r(M|\mathbf{y}) \propto \int_{\boldsymbol{\xi}_M} f_M(\mathbf{y}, \boldsymbol{\xi}_M) J_M(\mathbf{y}, \boldsymbol{\xi}_M) h_M(\boldsymbol{\xi}_M) d\boldsymbol{\xi}_M,$$

$$h_M(\boldsymbol{\xi}_M) = \begin{cases} 0 & \text{a smaller model predicts nearly as well} \\ 1 & \text{otherwise} \end{cases}$$

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- ▶ Motivated by non-local priors of [Johnson & Rossell \(2009\)](#)

Regression

- ▶ $\mathbf{Y} = \mathbf{X}\beta + \sigma\mathbf{Z}$
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- ▶ Better:

$$h_M(\beta_M) := I_{\{\frac{1}{2}\|\mathbf{X}^T(\mathbf{X}_M\beta_M - \mathbf{X}b_{min})\|_2^2 \geq \epsilon(n, |M|)\}}$$

where b_{min} solves

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- ▶ algorithm – [Bertsimas et al \(2016\)](#)
- ▶ similar to Dantzig selector [Candes & Tao \(2007\)](#)
different norm and target

GFD

$$r(M|\mathbf{y}) \propto \pi^{\frac{|M|}{2}} \Gamma\left(\frac{n - |M|}{2}\right) RSS_M^{-\left(\frac{n - |M| - 1}{2}\right)} E[h_M^\varepsilon(\beta_M^*)]$$

Observations:

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- ▶ $r(M|\mathbf{y})$ negligibly small for large models because of h , e.g., $|M| > n$ implies $r(M|\mathbf{y}) = 0$.
- ▶ Implemented using Grouped Independence Metropolis Hastings (Andrieu & Roberts, 2009).

Main Result

Theorem **Williams & H (2017+)**

Suppose the true model is given by M_T . Then under certain conditions, for a fixed positive constant $\alpha < 1$,

$$r(M_T|y) = \frac{r(M_T|y)}{\sum_{j=1}^{n^\alpha} \sum_{M:|M|=j} r(M|y)} \xrightarrow{P} 1 \text{ as } n, p \rightarrow \infty.$$

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- ▶ For a large model $|M| > p_T$ and large enough n or p ,

$$\frac{9}{2} \|X^T(H_M - H_{M(-1)})\mu_T\|_2^2 < \varepsilon_M(n, p),$$

where H_M and $H_{M(-1)}$ are the projection matrix for M and M with a covariate removed respectively.

Simulation

- ▶ Setup from [Rockova & George \(2015\)](#)
 - ▶ $n = 100, p = 1000, p_T = 8$.
 - ▶ Columns of X either a) independent or b) correlated with $\rho = 0.6$
 - ▶ $\varepsilon_M(n, p) = \Lambda_M \widehat{\sigma}_M^2 \left(\frac{n^{0.51}}{9} + |M| \frac{\log(p\pi)^{1.1}}{9} - \log(n)^\gamma \right)_+$ with $\gamma = 1.45$.

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 - ▶ Conditions of Theorem violated
 - ▶ based on conditions: p decreased to 500 to satisfy, performance improves.

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- ▶ What if small model not the right target, e.g., gene interactions?

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Recall

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- ▶ Conditioning U^* on $\{\mathbf{x} = \mathbf{G}(\mathbf{U}^*, \xi)\}$
 - “regularization by model”

Most general iid model

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There is a solution iff order of U_i^* matches order of x_i .

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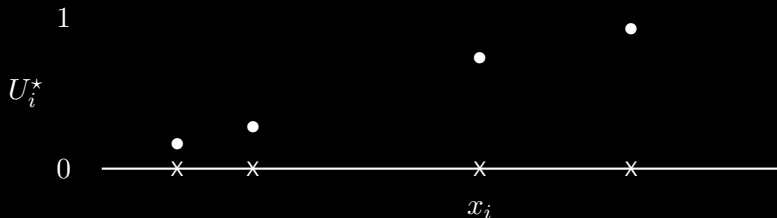
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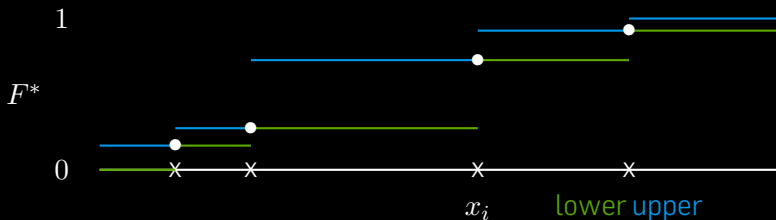
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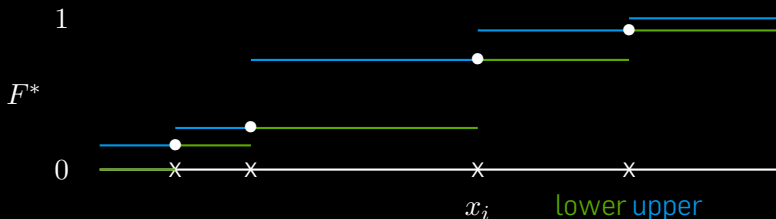
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- ▶ See Yifan Cui's poster for extension to censored data.

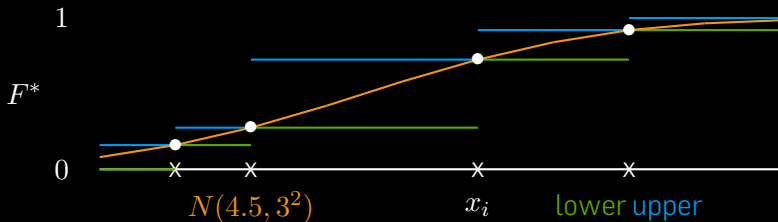
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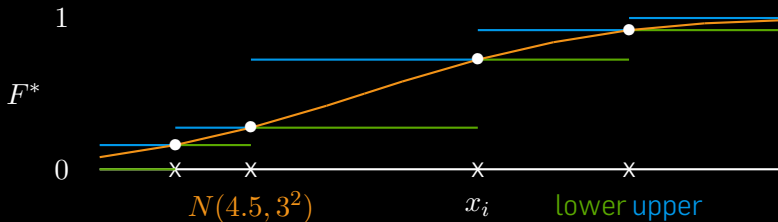
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- ▶ GFD is $r(\mu, \sigma) \propto \sigma^{-1} \prod_{i=1}^n \sigma^{-1} f(\sigma^{-1}(x_i - \mu))$

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- ▶ Option 1: condition $Z^*|x - Z^* > 0$
 - ▶ $r(\mu) = \frac{\varphi(x-\mu)}{\Phi(x)} I_{\{\mu>0\}}$
 - ▶ Lower confidence bounds do not have correct coverage.

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- ▶ Option 3: mixture
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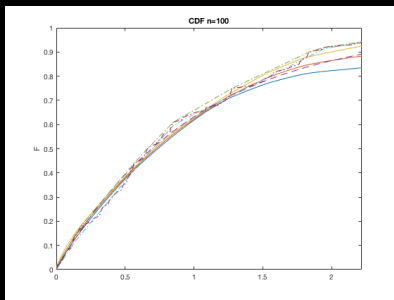
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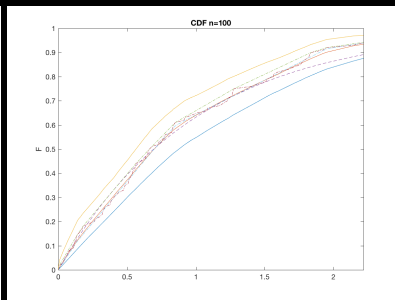
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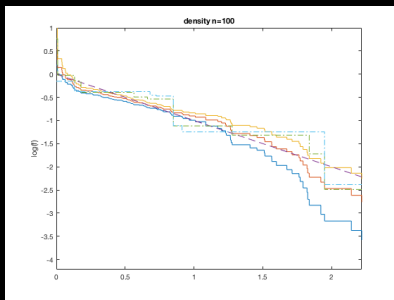
Condition



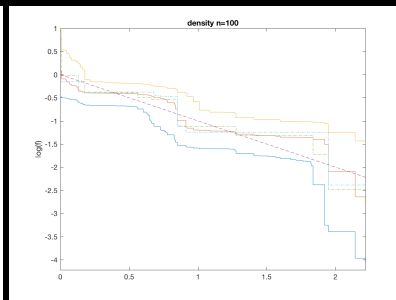
Projection

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Outline

- Introduction
- Definition
- Sparsity
- Regularization
- Conclusions

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- ▶ Applications
 - ▶ The proof is in the pudding

List of successful applications

- ▶ General Linear Mixed Models [E, H & Iyer \(2008\)](#); [Cisewski & H \(2012\)](#)
- ▶ Confidence sets for wavelet regression [H & Lee \(2009\)](#) and free knot splines [Sonderegger & H \(2014\)](#)
- ▶ Extreme value data (Generalized Pareto), Maximum mean, and model comparison [Wandler & H \(2011, 2012ab\)](#)
- ▶ Uncertainty quantification for ultra high dimensional regression [Lai, H & Lee \(2015\)](#), [Wandler & H \(2017+\)](#)
- ▶ Volatility estimation for high frequency data [Katsoridas & H \(2016+\)](#)
- ▶ Logistic regression with random effects (response models) [Liu & H \(2016,2017\)](#)

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Thank you!